# AN INCIDENTAL CATCH MODEL FOR PORBEAGLE ASSESSMENT AND STATUS EVALUATION 

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#### Abstract

SUMMARY Fisheries landings and associated biological data collection for porbeagle shark (Lamna nasus) declined substantially following CITES Appendix II trade restrictions in 2013 due to changes in fleet behavior. This document describes a new stock assessment method that can be used when length-frequency data and CPUE series are not available or reliable for index changes in abundance. The Incidental Catch Model (ICM) is based on the same general premise as datapoor, length-based assessments, in that it uses life history information and equilibrium assumptions to derive a theoretical age-structured population in the absence of fishing. In the ICM, the effect of historical fishing pressure on productivity is taken into account prior to evaluating fishery removals and abundance relative to reference points. The Northwest Atlantic stock was used to demonstrate the method, which can be adapted to assess stocks in the Northeast and South Atlantic by changing life history inputs and the removals series.


#### Abstract

RÉSUMÉ Les débarquements des pêcheries et la collecte de données biologiques associées concernant le requin-taupe commun (Lamna nasus) ont considérablement diminué suite aux restrictions commerciales découlant de l'inscription de cette espèce à l'annexe II de la CITES en 2013 en raison de changements dans le comportement des flottilles. Ce document décrit une nouvelle méthode d'évaluation des stocks qui peut être utilisée lorsque les données de fréquences de tailles et les séries de CPUE ne sont pas disponibles ou fiables pour les changements d'indice d'abondance. Le modèle de capture accidentelle («ICM» selon les sigles anglais) est basé sur le même principe général que les évaluations basées sur la taille et pauvres en données, dans la mesure où il utilise des informations sur le cycle vital et des postulats en conditions d'équilibre pour dériver une population théorique structurée par âge en l'absence de pêche. Dans le ICM, l'effet de la pression historique de la pêche sur la productivité est pris en compte avant d'évaluer les prises et l'abondance de la pêche par rapport à des points de référence. Le stock de l'Atlantique Nord-Ouest a été utilisé pour démontrer la méthode, qui peut être adaptée pour évaluer les stocks de l'Atlantique Nord-Est et Sud en modifiant les intrants du cycle vital et les séries de capture.


#### Abstract

RESUMEN Los desembarques de las pesquerías y la recopilación de datos biológicos asociados al marrajo sardinero (Lamna nasus) disminuyeron sustancialmente tras las restricciones al comercio del Apéndice II de la CITES en 2013 debido a los cambios en el comportamiento de la flota. En el presente documento se describe un nuevo método de evaluación de stock que puede utilizarse cuando no se dispone de datos de frecuencias de talla y series de CPUE o estos no son fiables para los cambios del índice de abundancia. El modelo de captura incidental (ICM) se basa en la misma premisa general que las evaluaciones basadas en la talla y con pocos datos, en el sentido de que utiliza la información del ciclo vital y los supuestos de equilibrio para derivar una población teórica estructurada por edades en ausencia de pesca. En el ICM se tiene en cuenta el efecto de la presión pesquera histórica en la productividad antes de evaluar las extracciones de la pesca y la abundancia con respecto a los puntos de referencia. Se utilizó el stock del Atlántico noroccidental para demostrar el método, que puede adaptarse para evaluar los stocks en el Atlántico nororiental y meridional cambiando las entradas del ciclo vital y las series de extracciones.


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## KEYWORDS

## Alternative assessment approaches, Incidental Catch Model, Porbeagle Shark, status

## 1. Introduction

A fundamental assumption underlying traditional fisheries assessment is that at least one of the data inputs (e.g. catch per unit effort; CPUE) indexes abundance. In situations where there have been profound changes in data indices that are independent of changes to abundance, such models (e.g. Surplus Production, Statistical Catch at Age, Virtual Population Analysis) will not provide information on population size or status relative to current fishery removals (Maunder et al. 2006). Similarly, data-poor approaches like length-based assessments rely on having a representative length-frequency distribution from the catches to compare with the equilibrium length distribution of the population (based on life history characteristics) to estimate fishing mortality (Hordyk et al. 2015a; Hordyk et al. 2015b). For the porbeagle stocks assessed by ICCAT, recent changes to CPUE series and relatively limited data from which to estimate length-frequency means that a new approach to stock assessment is needed.

Here, we propose an Incidental Catch Model (ICM) that can be used to evaluate historical abundance, estimate reference points, determine status (i.e. overfished), and quantify stock response to various levels of future fishing mortality. It incorporates variability in life history inputs (Cortés and Semba 2020) and enables status assessment relative to abundance reference points (i.e. overfished status; Brooks et al. 2010). We demonstrate the model by fitting it to data from the Northwest Atlantic porbeagle stock. We chose this example because the predicted abundance trajectory from the ICM could be directly compared to predictions from a fully-integrated Statistical Catch at Age model (Campana et al. 2010) to help with validation of the approach.

## 2. Methods

The ICM is a simulation model with two main parts: (1) a backward-projecting component, used to predict the historical abundance trajectory given the actual time series of removals, and (2) a forward-projecting component that accounts for any reduction in productivity from fishing pressure to assess current status as well as the scope for future removals. The ICM accounts for uncertainty in our understanding of porbeagle life history by simulating over a distribution of values for population productivity. This distribution comes from a Leslie matrix approach to get the theoretical maximum capacity for population growth in the absence of fishing ( $r_{\max }$ ). This type of modeling approach has been previously applied to demographic analyses of cetaceans (Caswell et al. 1998; Dans et al. 2003), basking shark (Campana et al. 2008) and white shark (Bowlby and Gibson 2020). This is the first application of this type of simulation model for stock assessment at ICCAT.

### 2.1 Simulation Model

It is very likely that porbeagle in the Northwest Atlantic are at low abundance, so we used a simple exponential growth model in the backwards projections to predict changes in population size $(N)$ from the current year ( $y$ ) to the previous year $(y-1)$, accounting for removals $(R)$ :
(1) $N_{y}=e^{r} N_{y-1}-R_{y-1}$
which can be rearranged as:
(2) $N_{y-1}=\frac{\left(N_{y}+R_{y-1}\right)}{e^{r}}$

The time series of removals from Task 1 data informs $R$, spanning the years 1961 to 2018 for the Northwest Atlantic. Thus, the backwards projections start in 2018 and end in 1961. The removals series incorporated an estimation of recent catches from non-reporting fleets based on a ratio estimator, consistent with the previous assessment (Anon 2010). We used the combined size distribution for captures (Santos et al. 2020) and the growth function parameters (Cortés and Semba 2020) for the Northwest Atlantic to transform Task 1 biomass into numbers.

The backwards projections occur on an annual time step where population increase happens first and removals take place afterwards. This makes the analysis more precautionary as it slightly exaggerates the effects of removals by modeling them as a discrete rather than a continuous process. Note that we used the population's capacity for growth in the absence of fishing $\left(r_{\max }\right)$ in these backwards projections. This means that the effect of removals in each year (i.e. the level of exploitation) was calculated relative to the population's theoretical capacity for growth at equilibrium, which is the same premise that underlies length-based assessment approaches (Hordyk et al. 2015a; Hordyk et al. 2015b).

In an age-structured population, fishing mortality $(F)$ and exploitation rates $(u)$ are a function of survival at age $\left(l_{x}\right)$, fisheries selectivity, removals and abundance. Survival at age is conditional on natural mortality $(M)$ and instantaneous fishing mortality $(F)$.
(3) $l_{x}=\prod_{i=0}^{x-1} e\left(-\left(M_{i}+F_{i}\right)\right)$

In the backward projections as well as the assessment of status (below), we found the value for F using minimization of the sum of squared residuals between observed removals ( $R_{y}$ ) and predicted removals. From the basic relationship between an annual exploitation rate $(u)$ and instantaneous fishing mortality $(F)$ :
(4) $u=1-e^{-F}$

The number of animals in the population in a given year $\left(N_{y}\right)$ that are vulnerable to the fishery becomes:
(5) vulnerable $=N_{y} \frac{\sum_{x=\text { sel }}^{A} l x}{\sum_{x=0}^{A} l x}$
where sel is the age at which animals are fully selected to the fishery and $A$ represents maximum age. Here we have approximated a dome-shaped selectivity function by assuming constant fishing mortality rates on juveniles and no fishing mortality on adults. This reflects the length composition data from various fleets, where the vast majority of fisheries captures are immature. We set $F$ to zero above the age at maturity in the calculation of ( $l_{x}$ ) and $\operatorname{sel}=1$. Predicted removals are simply vulnerable*u.

We used a simple logistic growth model for the forward projections of the simulation, assuming a carrying capacity $(K)$ of 20 million animals (i.e. very weak density dependence) in the Northwest Atlantic.
(6) $N_{t+1}=e^{r} N_{t}\left(1-\frac{N_{t}}{K}\right)$
(7)

This was done to ensure that the future projections could not grow without bound and thus produce overly optimistic results on the potential for population recovery. We assumed the same selectivity pattern as in the backwards projections and we accounted for the reduction to $r_{\max }$ from fisheries removals ( $r<r_{\max }$ ) by substituting the estimated value of $F$ into Equation 3. We projected the model forward for 50 years which is roughly 2.5 generation times for porbeagle (generation time $\sim 20$ years). Average removals from 2016-2018 were assumed for removals in 2019 and 2020, with the projections starting in 2021. This reflects the earliest year in which management changes could affect the population, given the lag between the 2020 assessment and the implementation of management recommendations.

### 2.2 Model parameterization

We obtained life history parameters for females from previously published data (refer to Table $\mathbf{1}$ in Cortés and Semba 2020). The growth equation came from Natanson et al. (2002) and information on reproductive biology, including the shape of the maturity ogive, litter size, and gestation period came from Jensen et al. (2002). The maturity ogive was updated to include unpublished data from Canada (Cortés and Semba 2020). An estimate of natural mortality $(M)$ is required to determine survival at age. We calculated six different estimators for $M$ (see details in Cortés and Semba 2020) and took the lowest of the suite in a particular iteration to calculate survivorship. Using the life history parameters, we applied the age-structured Leslie matrix approach (equivalent to a life table analysis/Euler-Lotka equation) for females described in Cortés and Semba (2020) to generate a distribution of $r_{\text {max }}$ values for input into the ICM. In brief, we used Monte Carlo sampling from predefined statistical distributions for multiple life history parameters to generate the maturity ogive and survivorship at age, as well as values for age at maturity and lifespan. The number of MC samples matched the total number of simulations run for the ICM ( $\mathrm{n}=$ 5000). Random draws for parameters contributing to the von Bertalanffy Growth Function (VBGF) and maturity ogive came from a multivariate normal distribution to account for covariance. We used a uniform distribution for lifespan with the observed longevity from vertebral aging as the lower bound and seven half-lives (from the $K$ parameter of the growth function) as the upper bound (Cortés and Semba 2020).

### 2.3 Status evaluation

We assessed whether the stock in the Northwest Atlantic is currently overfished using the SPR $_{\text {MER }}$ reference point proposed by Brooks et al. (2010): the Spawning Potential Ratio at Maximum Excess Recruitment. This biological reference point is derived entirely from life history data and has been found to accurately predict overfished status relative to Maximum Sustainable Yield (MSY) reference points from traditional stock assessments (Cortés and Brooks 2018). It is calculated as:
(8) $S P R_{M E R}=\frac{1}{\sqrt{\hat{\alpha}}}$

Where $\widehat{\boldsymbol{\alpha}}$ represents the maximum lifetime reproductive (MLR) rate (Myers et al. 1997, 1999). For this model the MLR is defined as the maximum number of female spawners that can be produced by a female spawner throughout her life (as in Bowlby \& Gibson 2020). It is calculated from the net reproductive rate or spawners per recruit (SPR) multiplied by maximum age-0 survival (Brooks et al. 2010). Overfished status is determined by comparing current abundance with a threshold value. This value typically represents a given proportion $(p)$ of the stock size which is expected to produce MSY (e.g. $p=1-M$; Brooks et al. 2010). However, ICCAT typically uses MSY directly as the reference point for abundance, which means that the threshold value is $p=1$.

The threshold value representing the depletion of spawners and recruits at Maximum Excess Recruitment, assuming a Beverton-Holt stock-recruit relationship, is:
(9) $\frac{s_{M E R}}{s_{0}}=\frac{\sqrt{\hat{\alpha}}-1}{\widehat{\alpha}-1}$

The population is considered overfished if the level of depletion in an abundance index $(I)$ divided by the threshold value in Equation (8) is smaller than the proportion $p$ :

$$
\text { (10) } \frac{\frac{I_{\text {current }}}{I_{\text {unfished }}}}{\frac{S_{M E R}}{S_{0}}}<\boldsymbol{p}
$$

Here, we have assumed that predicted abundance in 1961 represents unfished population size and that abundance in 2018 represents current abundance. Each iteration of the simulation yields a different value for $\widehat{\boldsymbol{\alpha}}$ as well as for current (2018) and unfished population size (1961) due to the manner in which variability is incorporated into the model (MC sampling from distributions). Therefore, solving Equation 8 gives a distribution of values that can be compared to $p$.

Similarly, the proportion of simulations that are overfished at a given time step in the forward projections can be found by using predicted future abundance as $\boldsymbol{I}_{\text {current }}$. We evaluated the forward predictions at 5 -year intervals relative to removals scenarios ranging from 0 to 24,000 animals $(0-734 \mathrm{mt})$. In each future year, the proportion of trajectories that are overfished becomes the number of simulations $<p$ divided by the total number of simulations.

### 2.4 Validation

There are no model residuals or deviances output by the ICM that can be used for validation with traditional metrics like AIC or negative loglikelihood (Johnson and Omland 2004). Instead, using different life history parameters or starting the population at a different initial abundance will lead to a different historical abundance trajectory and different estimates of $F$. For this reason, we felt it was necessary to compare model output with a more traditional fishery model to determine if the ICM was a reasonable assessment method. Fortunately, a Statistical Catch at Age (SCA) model that used data up to 2009 was developed for porbeagle in the Northwest Atlantic (Campana et al. 2010). The Canadian SCA was a fully integrated assessment, which incorporated the standardization of raw CPUE data within the model. It was sex-based and stage-structured, and partitioned catches according to three different age-length keys, representing catch profiles from three geographic regions. Four model variants were tested, where each incorporated slightly different life history assumptions leading to differences in productivity. Fisheries selectivity for each region was estimated in the model. Although 10 years old, the general structure of the SCA as well as the methods used for fitting (minimization of an objective function via the Laplace approximation in ADMB ) are still considered to be among best-practices for data-rich, age-structured assessments (Maunder and Piner 2015).

We qualitatively compared the predicted historical abundance trajectory from the ICM with predictions from the SCA, after standardizing data inputs into the ICM from the SCA. The SCA used a time series of removals derived from NAFO reporting (1961-2009) rather than Task 1 data, which was used as the removals series in the ICM for this comparison. The SCA assumed a Beverton-Holt stock recruitment relationship, and the variant called Model 3 had the most similar reproductive assumptions to the ICM. The VBGF growth parameters were identical.

### 2.5 Sensitivity Analyses

Recent research has suggested a possible biennial reproductive cycle for porbeagle (Natanson et al. 2019). Thus, we present two alternate life history parameterizations: one assuming a reproductive periodicity of one year (High Productivity), and one assuming a $50: 50 \mathrm{mix}$ of annually-reproducing and biennially-reproducing females (Medium Productivity). The Medium scenario is expected to better characterize productivity in the North Atlantic stock and was used as the base case in the current assessment. The High scenario is given both for continuity with the previous assessment (Anon 2010) and as a sensitivity analysis.

We also evaluated the sensitivity of the model to the assumed population size in 2018. The Canadian SCA model that was used for validation predicted an abundance of approximately 200,000 individuals in 2009, which was used as median abundance in 2018 in the base case of the ICM. This assumption was made because the ICM produced similar trends in abundance (i.e. similar overall decline rates) as compared to the SCA from this starting size, even though the absolute values predicted for abundance in each year differed. The ICM used a different removals series and had different productivity assumptions, so we would not expect an exact match in predicted abundance between the two models, yet we would expect similar trends. However, to evaluate a higher biomass scenario (representing a stock that was currently at a larger total size), we scaled the historical abundance trajectory so that the ICM predicted a population size of approximately 200,000 animals in 2009 for this sensitivity analysis. This was done by iteratively increasing assumed abundance in 2018 until median abundance in 2009 was 200,000 animals in the backwards projections.

## 3. Results

### 3.1 Base Case: mixture of biennial and annual reproduction

The median and $95^{\text {th }}$ percentiles for $r_{\text {max }}$ and the SPR $_{\text {MER }}$ reference point were $0.059(0.036,0.081)$ and 0.558 (0.415-0.743), respectively (Table 1). The backwards projections of the ICM indicated that maximum abundance occurred in the 1960s, with a median predicted population size of $\sim 985,000$ animals (top panel; Figure 1). Declines continued until 2010, reaching a low of $\sim 170,000$, before increasing to the assumed 2018 median population size of 200,000 animals. This means that the stock declined by $79 \%$ from maximum historical abundance over the entire time series. The threshold value for biomass at SPR $_{\text {MER }}$ was 353,000 animals.

The stock was predicted to be overfished with a very high probability, with $99 \%$ of iterations below the critical value of $p=1$. From the future projections (top panel; Figure 2), the stock would remain above the overfished threshold with a $>60 \%$ probability until 2070 if removals remained at or below 7,000 animals (Table 2). Even if removals (landings plus dead discards) were reduced to zero, it would take until 2035 to rebuild the population to the threshold level for biomass with $\mathrm{a}>60 \%$ probability.

### 3.2 Sensitivity Analysis: annual reproduction

If population productivity was higher, the median and $95^{\text {th }}$ percentiles for $r_{\text {max }}$ and the $\mathrm{SPR}_{\text {MER }}$ reference point were $0.079(0.057,0.102)$ and $0.456(0.339-0.607)$, respectively (Table 1). Predicted population size in 1961 was lower, at 849,000 animals (middle panel; Figure 1), leading to a predicted $76 \%$ decline over the entire time series. As would be expected from a population with higher productivity, biomass at $\mathrm{SPR}_{\text {MER }}$ was lower, at 266,000 animals. If reproduction is annual, the stock was still predicted to be overfished with a high probability, with $83 \%$ of iterations below the critical value of $p=1$. From the future projections (middle panel; Figure 2), the stock would remain above the overfished threshold with a $>60 \%$ probability for 2.5 generations if removals remained at or below 13,000 animals (Table 3). Note that it takes less time for the population to recover given that 2018 abundance is much closer to the overfished threshold (middle panel; Figure 2).

### 3.3 Sensitivity Analysis: higher biomass

This sensitivity run used the same life history parameters as in the base case, and had the same values for $r_{\max }$ and SPR $_{\text {MER }}$ (Table 1). Median population size in 2018 needed to be increased to 310,000 animals to give a predicted median abundance in 2009 of 200,000 animals. Population size in 1961 was 994,000 animals (lower panel; Figure 1), giving a historical decline rate of $68 \%$. The threshold value for biomass at $\operatorname{SPR}_{\text {MER }}$ was 349,000 animals. The slight difference in the threshold value from the base case arises from the random MC sampling.

The stock remained overfished if 2018 biomass was 310,000 animals, with $70 \%$ of iterations below the critical value of $p=1$. The future projections (lower panel; Figure 2) gave similar results to the high productivity sensitivity analysis; the stock would remain above the overfished threshold with a $>60 \%$ probability for 2.5 generations if removals were kept below 14,000 animals (Table 4). Similar to the previous scenario, the stock was close to the overfished threshold in 2018 ad could recover more quickly than in the base case scenario (lower panel; Figure 2)

### 3.4 Comparison of Base Case and Sensitivity Analyses

A mix of biennial and annual reproduction reduced the productive capacity of the stock by $25 \%$ as compared to annual reproduction. The SPR $_{\text {MER }}$ reference point increased from 0.456 ( $0.339-0.607$ ) to 0.558 ( $0.415-0.743$ ). Recall that the closer the SPR $_{\text {MER }}$ value is to 1 , the less exploitation can be permitted (Brooks et al. 2010).

The predicted trends in abundance are qualitatively similar among the base case and the two sensitivity analyses (Figure 1), but predicted abundance in 1961 was $\sim 136,000$ animals less when productivity was higher. This was expected because the population had greater ability to recover from exploitation in the High Productivity scenario so a smaller total number can support a given level of removals.

In all scenarios, the stock reached minimum abundance around 2002 and was predicted to have been increasing since that time. However, all scenarios suggest that the stock is currently overfished and that future removals need to remain low to permit population recovery.

### 3.5 Model Validation

When productivity inputs and the removals series used in the ICM were matched (as closely as was possible) to the Canadian SCA model, the different approaches produced extremely similar historical abundance trajectories. This could only occur if the ICM was producing similar estimates of the annual exploitation rate as compared to the SCA. Predicted starting population size in the 1960 s was just over 800,000 animals from both modeling approaches and declined following a very similar pattern (Figure 3). The total predicted decline rate was essentially identical. These results are quite remarkable given how different the data inputs, estimation method and assumptions are in the SCA as compared to the ICM and lends validity to our assertion that the ICM is useful for assessment.

## 4. Discussion and Conclusions

We have presented a relatively simple simulation model that can be used to assess status relative to biological reference points for porbeagle in the Northwest Atlantic. As in traditional stock assessment, the ICM assumes that the time series of removals is known. However, unlike traditional fisheries assessments, the ICM uses the theoretical productive capacity of the stock at equilibrium to index changes in abundance and to scale historical population size estimates. There was very good correspondence between the abundance predictions of the ICM as compared to a complex, sex-specific and age-structured SCA model. Because of this, we suggest that the ICM can be used to derive advice.

For the stock in the Northwest Atlantic, we used previous assessment results to scale 2018 abundance and to evaluate the sensitivity of the model to that assumption. In the absence of information on current population size, it would still be possible to use the ICM for assessment if a representative CPUE series was available. In this instance, the ICM would be parameterized so that relative changes in abundance matched the trends in the CPUE series (i.e. declined or increased by the same amount over the same number of years), giving a prediction of population size in each year of the removals series. We attempted to fit the ICM to the combined population in the South Atlantic by changing life history parameter input values (e.g. growth coefficients) and the removals series. However, the two CPUE series for the Southern stock showed very different trends over the same years, either
declining strongly in the SW or remaining essentially without trend in the SE. As in other types of fisheries models, the ICM was extremely sensitive to the opposite CPUE trends and the group decided it was not advisable to use the ICM to derive advice until a more representative CPUE series or an estimate of population size was available for the combined Southern stock.

It is very likely that porbeagle in the Northwest Atlantic are currently overfished and remain at low abundance. The future projections demonstrate that the stock in the Northwest Atlantic has very little capacity for removals. Keeping removals below 7,000 animals (base case) would convert to a biomass of 214 mt . For comparison, removals averaged 47 mt from 2014-2018, 143 mt from 2009-2013, 192 mt from 2004-2008 and 611 mt from 1999-2003. The ICM suggests that it is very unlikely that porbeagle in the Northwest Atlantic can sustain the level of removals seen prior to the early 2000s. Keeping current management restrictions in place will likely allow for porbeagle in the Northwest Atlantic to recover, even if current fisheries removals are slightly underestimated.

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Table 1. Demographic rates for the Northwest Atlantic population of porbeagle shark (Lamna nasus) used in the ICM model assuming a 50:50 mix of annual and biennial reproduction (Base Case) and the sensitivity run assuming annual reproduction (High Productivity).

|  | Base Case |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | Median | 0.025 | 0.975 | Median | 0.025 | 0.975 |
| Lambda | 1.061 | 1.037 | 1.084 | 1.083 | 1.059 | 1.107 |
| rmax | 0.059 | 0.036 | 0.081 | 0.079 | 0.057 | 0.102 |
| Generation Time | 20.116 | 17.317 | 21.324 | 19.317 | 16.948 | 20.462 |
| Net Repro Rate | 3.520 | 2.013 | 6.235 | 5.281 | 3.019 | 9.352 |
| Age-0 Survivorship | 0.912 | 0.894 | 0.931 | 0.912 | 0.894 | 0.931 |
| Alpha hat | 3.208 | 1.812 | 5.803 | 4.811 | 2.718 | 8.705 |
| Steepness | 0.445 | 0.312 | 0.592 | 0.546 | 0.405 | 0.685 |
| SPR_MER | 0.558 | 0.415 | 0.743 | 0.456 | 0.339 | 0.607 |
| R | 0.603 | 0.543 | 0.702 | 0.554 | 0.508 | 0.621 |

Table 2. Percent of simulations that are above the overfished threshold for porbeagle (Lamna nasus) in the Northwest Atlantic for the base case of the ICM.

| Removals (\#) | Removals (mt) | 2020 | 2025 | 2030 | 2035 | 2040 | 2045 | 2050 | 2055 | 2060 | 2065 | 2070 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 0 | 2 | 21 | 47 | 68 | 83 | 92 | 96 | 98 | 99 | 99 | 100 |
| 1000 | 31 | 3 | 21 | 44 | 63 | 77 | 87 | 92 | 95 | 97 | 98 | 99 |
| 2000 | 61 | 2 | 19 | 40 | 57 | 71 | 81 | 87 | 91 | 94 | 95 | 96 |
| 3000 | 92 | 1 | 16 | 35 | 50 | 62 | 72 | 79 | 85 | 88 | 90 | 92 |
| 4000 | 122 | 2 | 15 | 32 | 47 | 58 | 66 | 73 | 78 | 82 | 84 | 87 |
| 5000 | 153 | 2 | 13 | 27 | 41 | 50 | 58 | 64 | 68 | 72 | 76 | 78 |
| 6000 | 183 | 1 | 12 | 25 | 37 | 45 | 52 | 57 | 62 | 65 | 67 | 70 |
| 7000 | 214 | 2 | 10 | 22 | 32 | 39 | 46 | 50 | 54 | 57 | 60 | 62 |
| 8000 | 245 | 2 | 10 | 19 | 27 | 34 | 39 | 44 | 47 | 50 | 53 | 55 |
| 9000 | 275 | 2 | 8 | 17 | 23 | 30 | 34 | 38 | 41 | 43 | 45 | 47 |
| 10000 | 306 | 2 | 8 | 14 | 20 | 25 | 29 | 31 | 34 | 36 | 38 | 39 |
| 11000 | 336 | 1 | 6 | 13 | 17 | 21 | 25 | 27 | 29 | 31 | 32 | 33 |
| 12000 | 367 | 2 | 7 | 11 | 15 | 18 | 21 | 23 | 24 | 26 | 27 | 28 |
| 13000 | 398 | 2 | 5 | 9 | 12 | 14 | 16 | 18 | 19 | 20 | 21 | 22 |
| 14000 | 428 | 2 | 5 | 7 | 9 | 12 | 13 | 14 | 15 | 16 | 17 | 18 |
| 15000 | 459 | 1 | 3 | 5 | 6 | 8 | 9 | 10 | 11 | 11 | 12 | 12 |
| 16000 | 489 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 9 | 10 | 10 |
| 17000 | 520 | 2 | 2 | 3 | 4 | 5 | 5 | 6 | 6 | 6 | 7 | 7 |
| 18000 | 550 | 2 | 2 | 2 | 3 | 3 | 4 | 4 | 4 | 5 | 5 | 5 |
| 19000 | 581 | 2 | 1 | 2 | 2 | 3 | 3 | 3 | 3 | 3 | 3 | 4 |
| 20000 | 612 | 2 | 1 | 1 | 2 | 2 | 2 | 2 | 2 | 2 | 3 | 3 |
| 21000 | 642 | 2 | 1 | 1 | 1 | 1 | 1 | 2 | 2 | 2 | 2 | 2 |
| 22000 | 673 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 23000 | 703 | 2 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 24000 | 734 | 2 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

Table 3. Percent of simulations that are above the overfished threshold for the sensitivity analysis assuming annual reproduction (High Productivity) in the ICM.

| Removals <br> (\#) | Removals (mt) | 2020 | 2025 | 2030 | 2035 | 2040 | 2045 | 2050 | 2055 | 2060 | 2065 | 2070 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 0 | 34 | 67 | 91 | 99 | 100 | 100 | 100 | 100 | 100 | 100 | 100 |
| 1000 | 31 | 35 | 66 | 88 | 98 | 100 | 100 | 100 | 100 | 100 | 100 | 100 |
| 2000 | 61 | 35 | 65 | 85 | 96 | 99 | 100 | 100 | 100 | 100 | 100 | 100 |
| 3000 | 92 | 33 | 61 | 80 | 92 | 97 | 99 | 100 | 100 | 100 | 100 | 100 |
| 4000 | 122 | 35 | 60 | 78 | 89 | 94 | 97 | 99 | 99 | 100 | 100 | 100 |
| 5000 | 153 | 33 | 58 | 74 | 84 | 90 | 94 | 96 | 97 | 98 | 99 | 99 |
| 6000 | 183 | 34 | 56 | 70 | 80 | 86 | 90 | 93 | 94 | 95 | 96 | 97 |
| 7000 | 214 | 33 | 54 | 67 | 76 | 82 | 86 | 89 | 90 | 92 | 93 | 93 |
| 8000 | 245 | 34 | 53 | 64 | 72 | 77 | 81 | 84 | 86 | 87 | 89 | 89 |
| 9000 | 275 | 34 | 50 | 61 | 68 | 72 | 76 | 79 | 81 | 82 | 83 | 84 |
| 10000 | 306 | 34 | 49 | 58 | 64 | 68 | 71 | 73 | 75 | 76 | 77 | 78 |
| 11000 | 336 | 34 | 47 | 55 | 60 | 64 | 66 | 68 | 70 | 71 | 72 | 73 |
| 12000 | 367 | 35 | 46 | 53 | 57 | 60 | 62 | 63 | 65 | 66 | 67 | 67 |
| 13000 | 398 | 35 | 44 | 50 | 53 | 56 | 58 | 59 | 60 | 61 | 61 | 62 |
| 14000 | 428 | 35 | 42 | 47 | 50 | 52 | 53 | 54 | 55 | 56 | 56 | 57 |
| 15000 | 459 | 32 | 39 | 43 | 45 | 46 | 48 | 48 | 49 | 50 | 50 | 51 |
| 16000 | 489 | 33 | 37 | 40 | 42 | 43 | 44 | 44 | 45 | 45 | 45 | 45 |
| 17000 | 520 | 34 | 36 | 38 | 38 | 39 | 39 | 40 | 40 | 40 | 40 | 40 |
| 18000 | 550 | 33 | 34 | 34 | 34 | 34 | 35 | 35 | 35 | 35 | 35 | 35 |
| 19000 | 581 | 34 | 33 | 32 | 32 | 31 | 31 | 31 | 31 | 30 | 31 | 31 |
| 20000 | 612 | 34 | 31 | 30 | 29 | 28 | 28 | 27 | 27 | 27 | 27 | 26 |
| 21000 | 642 | 35 | 31 | 27 | 26 | 25 | 24 | 23 | 23 | 22 | 22 | 22 |
| 22000 | 673 | 33 | 28 | 24 | 22 | 21 | 20 | 19 | 19 | 18 | 18 | 18 |
| 23000 | 703 | 35 | 26 | 22 | 20 | 18 | 17 | 16 | 15 | 15 | 15 | 14 |
| 24000 | 734 | 34 | 24 | 20 | 16 | 15 | 14 | 13 | 12 | 12 | 11 | 11 |

Table 4. Proportion of simulations that are above the overfished threshold for the sensitivity analysis assuming a larger population size in 2009 in the ICM.

| Removals <br> (\#) | Removals (mt) | 2020 | 2025 | 2030 | 2035 | 2040 | 2045 | 2050 | 2055 | 2060 | 2065 | 2070 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 0 | 45 | 73 | 92 | 97 | 99 | 100 | 100 | 100 | 100 | 100 | 100 |
| 1000 | 31 | 52 | 77 | 93 | 97 | 99 | 100 | 100 | 100 | 100 | 100 | 100 |
| 2000 | 61 | 49 | 75 | 90 | 96 | 97 | 99 | 99 | 99 | 99 | 99 | 100 |
| 3000 | 92 | 50 | 72 | 87 | 95 | 96 | 98 | 99 | 99 | 100 | 100 | 100 |
| 4000 | 122 | 46 | 68 | 84 | 91 | 95 | 96 | 97 | 97 | 98 | 99 | 99 |
| 5000 | 153 | 47 | 68 | 80 | 90 | 94 | 95 | 96 | 97 | 97 | 98 | 98 |
| 6000 | 183 | 48 | 64 | 77 | 86 | 91 | 93 | 95 | 96 | 96 | 96 | 96 |
| 7000 | 214 | 44 | 61 | 73 | 80 | 85 | 88 | 91 | 92 | 93 | 94 | 94 |
| 8000 | 245 | 46 | 63 | 73 | 79 | 84 | 86 | 88 | 88 | 89 | 90 | 91 |
| 9000 | 275 | 38 | 51 | 62 | 68 | 74 | 79 | 80 | 82 | 84 | 85 | 85 |
| 10000 | 306 | 43 | 55 | 63 | 70 | 74 | 77 | 78 | 80 | 81 | 83 | 83 |
| 11000 | 336 | 48 | 59 | 66 | 69 | 72 | 75 | 76 | 77 | 77 | 78 | 79 |
| 12000 | 367 | 42 | 50 | 55 | 59 | 61 | 64 | 66 | 67 | 67 | 68 | 69 |
| 13000 | 398 | 45 | 51 | 54 | 58 | 60 | 62 | 63 | 65 | 65 | 65 | 66 |
| 14000 | 428 | 48 | 52 | 55 | 55 | 57 | 57 | 58 | 59 | 60 | 60 | 61 |
| 15000 | 459 | 50 | 52 | 52 | 53 | 53 | 54 | 53 | 52 | 52 | 53 | 54 |
| 16000 | 489 | 45 | 47 | 47 | 47 | 47 | 47 | 47 | 46 | 46 | 46 | 47 |
| 17000 | 520 | 46 | 44 | 44 | 43 | 43 | 43 | 42 | 42 | 41 | 42 | 41 |
| 18000 | 550 | 47 | 43 | 40 | 39 | 38 | 38 | 37 | 36 | 36 | 35 | 35 |
| 19000 | 581 | 44 | 39 | 37 | 37 | 35 | 33 | 33 | 32 | 31 | 32 | 32 |
| 20000 | 612 | 45 | 39 | 35 | 32 | 30 | 28 | 28 | 28 | 28 | 28 | 27 |
| 21000 | 642 | 47 | 42 | 36 | 32 | 30 | 28 | 27 | 25 | 25 | 24 | 24 |
| 22000 | 673 | 38 | 28 | 24 | 21 | 19 | 18 | 17 | 17 | 17 | 16 | 16 |
| 23000 | 703 | 43 | 30 | 24 | 21 | 18 | 16 | 15 | 15 | 14 | 13 | 13 |
| 24000 | 734 | 52 | 36 | 28 | 23 | 18 | 17 | 14 | 13 | 13 | 12 | 12 |



Figure 1. Predicted median abundance trajectory (solid line) and $80^{\text {th }}$ percentiles (dashed lines) for the porbeagle stock in the Northwest Atlantic from 1961 to 2018 for the base case of the ICM (top panel), the sensitivity run assuming higher productivity (middle panel) and the sensitivity run assuming higher abundance in 2009 (lower panel).





Median Removals
$\begin{array}{ll}-0 & -13000 \\ -1000 & -14000 \\ -2000 & -15000 \\ -3000 & -16000 \\ -4000 & -17000 \\ -5000 & -18000 \\ -6000 & -19000 \\ -7000 & -20000 \\ -8000 & -21000 \\ -9000 & -22000 \\ -10000 & -23000 \\ -11000 & -24000 \\ -12000 \\ \end{array}$

Figure 2. Predicted relative abundance for removals ranging from 0 to 24,000 animals (coloured lines), expressed as the biomass/biomass at SPR $_{\text {MER }}$ ratio for the base case of the ICM (top panel), the sensitivity run assuming high productivity (middle panel) and the sensitivity run assuming greater 2009 abundance (lower panel). The horizontal line shows the $B_{S P R_{M E R}}$ reference point and the projections extend for 50 years.


Figure 3. Comparison of the predicted historical abundance trajectories from a Statistical Catch at Age Model (Campana et al. 2010) for Porbeagle in the Northwest Atlantic (top panel) with output from the ICM over the same number of years, using the NAFO series for removals (lower panel).


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