# STANDARDIZED CATCH RATES OF BLUE SHARKS IN THE WESTERN NORTH ATLANTIC OCEAN FROM THE U.S. PELAGIC LONGLINE OBSERVER PROGRAM 1992-2014 AND 2015-2021 

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#### Abstract

SUMMARY An updated index of abundance was developed for blue shark (Prionace glauca) from data collected by the U.S. pelagic longline fishery observer program (1992-2021). The U.S. pelagic longline CPUE was estimated for two separate periods (1992-2014 and 2015-2021) to account for changes in U.S. management regulations that could not be included in the standardization process. The indices were calculated using a two-step delta-lognormal approach that treats the proportion of positive sets and the CPUE of positive catches separately. Observations affected by fishing regulations such as time-area closures or bait restrictions were excluded in this analysis. The standardized indices with $95 \%$ confidence intervals are reported. The standardized index 1992-2014 showed an initial increasing trend from 1992 to 1998, followed by a decrease to 2003, an increase to 2011, and a subsequent decrease to 2014. The standardized index 2015-2021 showed an overall decreasing trend.


## RÉSUMÉ

Un indice d'abondance actualisé a été développé pour le requin peau bleue (Prionace glauca) à partir des données collectées par le programme d'observateurs de la pêche pélagique palangrière des États-Unis (1992-2021). La CPUE des palangriers pélagiques américains a été estimée pour deux périodes distinctes (1992-2014 et 2015-2021) afin de tenir compte des changements dans les réglementations de gestion américaines qui n'ont pas pu être inclus dans le processus de standardisation. Les indices ont été calculés au moyen d'une approche deltalognormale en deux étapes qui traite séparément la proportion d'opérations positives et la CPUE des captures positives. Les observations affectées par des réglementations de pêche telles que les fermetures spatio-temporelles ou les restrictions sur les appâts ont été exclues de cette analyse. Les indices standardisés avec des intervalles de confiance à 95\% sont déclarés. L'indice standardisé de 1992-2014 présentait une tendance initiale à la hausse de 1992 à 1998, suivie d'une baisse jusqu'en 2003, d'une hausse jusqu'en 2011 et d'une baisse ultérieure jusqu'en 2014. L'indice standardisé 2015-2021 affichait une tendance globale à la baisse.

## RESUMEN

Se desarrolló un índice actualizado de abundancia para el tiburón azul (Prionace glauca) a partir de los datos recogidos por el programa de observadores de la pesquería de palangre pelágico de Estados Unidos (1992-2021). La CPUE de palangre pelágico de Estados Unidos se estimó para dos periodos diferentes (1992-2014 y 2015-2021) para tener en cuenta los cambios en los reglamentos de ordenación estadounidenses que no pudieron incluirse en el proceso de estandarización. Los índices se calcularon utilizando un enfoque delta-lognormal de dos etapas que trata la proporción de lances positivos y la CPUE de las capturas positivas por separado. En este análisis se excluyeron las observaciones afectadas por reglamentaciones pesqueras como las vedas espaciotemporales o las restricciones de cebo. Se comunican los índices estandarizados con intervalos de confianza del $95 \%$. El índice estandarizado 1992-2014 mostraba una tendencia inicial creciente de 1992 a 1998, seguida de un descenso hasta 2003,

[^0]
# un aumento hasta 2011 y un posterior descenso hasta 2014. El índice estandarizado 2015-2021 

 mostró una tendencia global decreciente.KEYWORDS<br>Catch/effort, Commercial fishing, Longlining, Pelagic fisheries, Shark fisheries, Observer programs, Blue shark

## 1. Introduction

A relative abundance index from the U.S. commercial pelagic longline fishery observer program was generated and used in the 2004, 2008, and 2015 ICCAT assessments of blue sharks (ICCAT 2005, 2009, 2016). In this document, the commercial series is updated to examine recent trends in abundance of blue sharks for use in the 2023 stock assessment of the North Atlantic stock. Indices of abundance for blue sharks from the U.S. commercial pelagic longline fishery observer program were previously developed by Brooks et al. (2005), Cortés (2007; 2009, 2016), and Cortés et al., (2007). The U.S. pelagic longline observer program managed by the U.S. NOAA/NMFS/Southeast Fisheries Science Center, currently has a target coverage of 8-percent of the longline sets deployed by the fleet. The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations affecting its pelagic longline fleet. The two indices were calculated using a two-step delta-lognormal approach that treats the proportion of positive sets and the CPUE of positive catches separately.

## 2. Materials and Methods

### 2.1 Data

The pelagic longline fishing grounds for the U.S. fleet traditionally extended from the Grand Banks in the North Atlantic to $5-10^{\circ}$ south, off the South American coast, including the Caribbean and the Gulf of Mexico. However, reductions in the number of vessels actively participating in the longline fishery has reduced the area of operation of the fleet during the past few years. Eleven U.S. domestic geographical areas of longline fishing have traditionally been used to analyze fisheries data (Figure 1): the Caribbean (CAR, area 1), Gulf of Mexico (GOM, area 2), Florida East coast (FEC, area 3), South Atlantic Bight (SAB, area 4), Mid-Atlantic Bight (MAB, area 5), New England coastal (NEC, area 6), Northeast distant waters (NED, or Grand Banks, area 7), Sargasso (SAR, area 8), North Central Atlantic (NCA, area 9), Tuna North (TUN, area 10), and Tuna South (TUN, area 11).

Data from the U.S. pelagic longline observer program were available for 1992-2021. The observer dataset was restricted to areas 5, 6 , and 7 (north of $35^{\circ}$ North latitude) owing to insufficient and unbalanced blue shark catch observations by year in the remaining areas. Areas 5, 6, and 7 accounted for 90 and $89 \%$ of all observations in the observer dataset for period 1992-2014 and period 2015-2021, respectively (Figure 2).

Several data restrictions were implemented in the present analysis to account for time-area closures or bait restrictions following Walter and Lauretta (2015) and Cortés (2017). Due to the different effects of spatio-temporal closures in different areas, a single "closure" effect could not be considered because it would likely differ among areas and thus the most parsimonious approach was to exclude data from the entire time series before and after the closure for each area. More specifically, the following data restrictions were applied: (1) the DeSoto Canyon Closed Area in the Gulf of Mexico, closed year-round; (2) the East Florida Coast Closed Area, closed year-round; (3) the Charleston Bump Closed Area, closed February-April; (4) the Northeastern United States Closed Area, closed in June; (5) the Northeast Distant Gear Restricted Area, closed year-round except for specific bait-gear configurations; and (6) the Spring Gulf of Mexico Gear Restricted Areas, closed April-May (Figure 1).

Based on the methodology used by Brooks et al. (2005), Cortés (2007, 2009, 2016), and Cortés et al. (2007), the following factors were considered in the analysis: year, area, quarter (January-March, April-June, July-September, October-December), presence or absence of light sticks, and whether or not the data were part of experimental fishing (conducted in years 2000-2003 in the Northeast Distant area only). Additionally, nominal catch rates (catch per thousand hooks) of swordfish, Xiphias gladius, and tuna (the sum of albacore, Thunnus alalunga, skipjack, Euthynnus pelamis, bigeye, Thunnus obesus, and yellowfin tuna, Thunnus albacares) were calculated for each set, and a categorical factor based on the quartile of those catch rates was assigned to each set (the factors are denoted
as Sqr and Tqr, respectively). The reason for creating these factors, which correspond to the $<25 \%, 25-49 \%$, $50-$ $75 \%$, and $>75 \%$ of the proportion, was to attempt to control for effects of blue shark catch rates associated with changes of fishing operations when the fleets target species changed. We also considered the following interactions: year*area, year*quarter, as well as the interactions between area and the nominal catch rate quartiles for tuna and swordfish (area*Sqr and area*Tqr). Blue shark nominal and standardized CPUEs were defined/estimated as total catch per 1000 hooks where total catch included all retained, dead discarded, and live released blue sharks.

### 2.2 Analysis

Relative abundance indices were estimated using a Generalized Linear Modeling (GLM) approach assuming a delta lognormal model distribution. A binomial error distribution is used for modeling the proportion of positive sets with a logit function as link between the linear factor component and the binomial error. A lognormal error distribution is used for modeling the catch rates of successful sets, wherein estimated CPUE rates assume a lognormal distribution (lnCPUE) of a linear function of fixed factors. The models were fitted with the SAS GENMOD procedure using a forward stepwise approach in which each potential factor was tested one at a time. Initially, a null model was run with no explanatory variables (factors). Factors were then entered one at a time and the results ranked from smallest to greatest reduction in deviance per degree of freedom when compared to the null model. The factor which resulted in the greatest reduction in deviance per degree of freedom was then incorporated into the model if two conditions were met: 1) the effect of the factor was significant at least at the 5\% level based on the results of a Chi-Square statistic of a Type III likelihood ratio test, and 2) the deviance per degree of freedom was reduced by at least $1 \%$ with respect to the less complex model. Single factors were incorporated first, followed by fixed first-level interactions. The year factor was always included because it is required for developing a time series. Results were summarized in the form of deviance analysis tables including the deviance for proportion of positive observations and the deviance for the positive catch rates.

Once the final model was selected, it was run using the SAS GLIMMIX macro (which itself uses iteratively reweighted likelihoods to fit generalized linear mixed models with the SAS MIXED procedure; Wolfinger and O'Connell 1993, Littell et al. 1996). In this model, any interactions that included the year factor were treated as a random effect. Goodness-of-fit criteria for the final model included Akaike's Information Criterion (AIC), Schwarz's Bayesian Criterion, and $-2^{*}$ the residual log likelihood ( -2 Res L). The significance of each individual factor was tested with a Type III test of fixed effects, which examines the significance of an effect with all the other effects in the model (SAS Institute Inc. 1999). The final mixed model calculated relative indices as the product of the year effect least squares means (LSMeans) from the binomial and lognormal components. LSMeans estimates were weighted proportionally to observed margins in the input data, and for the lognormal estimates, a back-transformed log bias correction was applied (Lo et al. 1992).

## 3. Results

### 3.1 Period 1 (1992-2014)

Factors retained for the blue shark proportion of positive sets were Sqr, year, Tqr, area, and quarter; and for the positive catches, the factors area, quarter, year, Sqr, Tqr, year*quarter, year*area, and Sqr *area were retained (Table 1a). The factor Sqr and the factor area explained $54.33 \%$ and $48.12 \%$ of the deviance for the proportion positive and positive catches, respectively (Table 2a). The estimated standardized CPUE and CV values are given in Table 3a. The nominal index showed an overall decreasing trend: specifically, a decreasing trend from 1992 to 2003, followed by an increase from 2003 to 2011, then followed by a decreasing tendency thereafter (Figure 3a). In contrast, the standardized index showed an initial increasing trend from 1992 to 1998, followed by a decrease to 2003, an increase to 2011, and a subsequent decrease to 2014. The sometimes large interannual fluctuations in the observer index may be due to small sample size (15-274 positive sets per year; Figure 3a). Diagnostic plots showed some patterns in the residuals of the proportion positive sets (Figure 4a).

### 3.2 Period 2 (2015-2021)

Factors retained for the blue shark proportion of positive sets were quarter, area, Tqr, and year; and for the positive catches, the factors year, quarter, area, year*quarter, Tqr*area, and Sqr*area were retained (Table 1b). The factor quarter and the factor year*quarter explained $44.59 \%$ and $25.41 \%$ of the deviance for the proportion positive and positive catches, respectively (Table 2b). The estimated standardized CPUE and CV values are given in Table 3b. The nominal index showed an overall decreasing trend: specifically, a decreasing trend from 2015 to 2019,
followed by an increase from 2019 to 2021 (Figure 3b). In contrast, the standardized index showed an initial increasing trend from 2015 to 2016, followed by a decrease to 2020, an increase to 2021. The sometimes large interannual fluctuations in the observer index may be due to small sample size ( $62-210$ positive sets per year; Figure 3b). Diagnostic plots showed some patterns in the residuals of the proportion positive sets (Figure 4b).

## 4. Discussion

Trends in relative abundance predicted from this analysis and a previous analysis (Cortés 2016) were similar during the overlapping years (1992-2013), with both series showing a concave shape, consisting of an initial decline to about the mid-2000s, followed by an increasing trend thereafter (Figure 3a). However, the proportion of positive sets in 2001, 2002 and 2003 was much higher in the current analysis. This discrepancy may due to recent changes in the observer database, but it did not affect the predicted trends in relative abundance (Figure 3a). The observer dataset has small sample sizes leading to large interannual variation. Despite the fact that blue sharks are likely the most productive pelagic shark species (Cortés and Brooks 2018), sharp interannual changes in abundance, such as those sometimes displayed by the observer series are inconsistent with the biology of most sharks, whose stock abundance is expected to fluctuate relatively little from year to year. Similar to the finding of mako sharks based on the same U.S. longline observer program (Cortés 2017), considering the time-area closures did not appear to overly influence the predicted catch rates. It is also unlikely that other management actions, such as quota reductions, may have had any effect on the catch rates of blue sharks because the pelagic longline fishery in the USA has not traditionally targeted them, and catch rates used here are based on total catch (the sum of blue sharks kept, discarded dead and released alive). Other factors, such as hook size and type, were not included in the analysis, but may have affected catch rates of blue sharks. Fishing depth was indirectly taken into account in our analysis by using proxies for fishers targeting swordfish or tunas, but we did not differentiate between different species of tunas being targeted. The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations. Up to 2014, U.S. pelagic longline vessels used to land blue sharks which allowed them to land a limited number of bluefin tuna. However, the adoption of a bluefin tuna IBQ (individual bycatch quota) for each individual vessel in the U.S. longline fleet allowed them to land bluefin tuna without the requirement of landing a certain required amount of other target species. Since there is absolutely no commercial value in catching blue sharks for the U.S. pelagic longline fleet, the fleet gradually changed their fishing behavior to avoid areas with large blue shark catches. Therefore, changes in the proportion of positives due to the changes in U.S. management regulations warranted the estimation of 2 separate CPUE series. However, this splitting does not seem to have affected the overall trend of this relative abundance index (Table 4 and Figure 5).

## References

Brooks, E.N., M. Ortiz, L.K. Beerkircher, and Y. Apostolaki. 2005. Standardized catch rates for blue shark and shortfin mako shark from the U.S. pelagic logbook and U.S. pelagic observer program, and U.S. weighout data. Collect. Vol. Sci. Pap. ICCAT, 58(3): 1054-1072.

Cortés, E. 2007. Standardized catch rates for blue shark and shortfin mako shark from the US pelagic longline logbook and observer programs. Collect. Vol. Sci. Pap. ICCAT, 60(2): 617-628.

Cortés, E. 2009. Standardized catch rates for blue shark and mako sharks from the US pelagic longline logbook and observer programs. Collect. Vol. Sci. Pap. ICCAT, 64(5): 1595-1613.

Cortés, E. 2016. Standardized catch rates of blue sharks in the western North Atlantic Ocean from the US pelagic longline logbook and observer programs. Collect. Vol. Sci. Pap. ICCAT, 72(4): 1067-1082.

Cortés, E. 2017. Stock status indicators of mako sharks in the western North Atlantic Ocean based on the US pelagic longline logbook and observer programs. Collect. Vol. Sci. Pap. ICCAT, 73(8): 2891-2910.

Cortés, E., C.A. Brown, and L.K. Beerkircher. 2007. Relative abundance of pelagic sharks in the western North Atlantic Ocean, including the Gulf of Mexico and Caribbean Sea. Gulf and Caribbean Research 19: 37-52.

Cortés, E., and E.N. Brooks. 2018. Stock status and reference points for sharks using data-limited methods and life history. Fish and Fisheries 19: 1110-1129.

ICCAT 2005. Report of the 2004 Inter-sessional meeting of the ICCAT sub-committee on by-catches: shark stock assessment. Collect. Vol. Sci. Pap. ICCAT, 58(3): 799-890.

ICCAT 2009. Report of the 2008 Shark stock assessments Meeting. Collect. Vol. Sci. Pap. ICCAT, 64(5): 13431491.

ICCAT 2016. Report of the 2015 Shark stock assessments Session. Collect. Vol. Sci. Pap. ICCAT, 72(4): 8661019.

ICCAT 2017. Report of the 2017ICCAT Shortfin Mako assessments Meeting. Collect. Vol. Sci. Pap. ICCAT, 74(4): 1465-15611.

Littell, R.C., G.A. Milliken, W.W. Stroup, and R.D Wolfinger. 1996. SAS® System for Mixed Models. Cary, NC: SAS Institute Inc., 1996. 663 pp.

Lo, N.C., L.D. Jacobson, and J.L. Squire. 1992. Indices of relative abundance from fish spotter data based on delta-lognormal models. Can. J. Fish. Aquat. Sci. 49: 2515-2526.

SAS Institute, INC. 1999. SAS/STAT User's Guide, version 8, NC: SAS Institute Inc., 1999. 3884 pp.
Walter, J. and M. Lauretta. Standardized catch rates for bigeye tune (Thunnus obesus) from the United States pelagic longline fishery. SCRS/2015/082.

Wolfinger, R. and M. O’Connell. 1993. Generalized linear mixed models: a pseudo-likelihood approach. J. Stat. Comput. Simul. 48: 233-243.

Table 1a. Factors retained in the model of proportion of positive sets and positive catch of blue sharks for the U.S. pelagic longline observer program data (1992-2014). The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.


Table 1b. Factors retained in the model of proportion of positive sets and positive catch of blue sharks for the U.S. pelagic longline observer program data (2015-2021). The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.

| Model factors proportion positives | DF | Deviance | Deviance/DF | \%reduction | Log-likehood |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Null | 1625 | 2201 | 1.355 |  | -1101 |
| Quarter | 1622 | 2020 | 1.245 | 8.08\% | -1010 |
| Quarter Area | 1621 | 1895 | 1.169 | 6.15\% | -947 |
| Quarter Area Tqr | 1618 | 1840 | 1.138 | 2.67\% | -920 |
| Final model: Quarter Area Tqr Year | 1612 | 1794 | 1.113 | 2.14\% | -897 |
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|  |  |  |  |  |  |
|  |  |  |  |  |  |
|  |  |  |  |  |  |
| Model factors positive catch rates | DF | Deviance | Deviance/DF | \%reduction | Log-likehood |
|  |  |  |  |  |  |
| Null | 958 | 1116 | 1.165 |  | -1433 |
| Year | 952 | 1060 | 1.114 | 4.37\% | -1409 |
| Year Quarter | 949 | 1023 | 1.078 | 3.25\% | -1392 |
| Year Quarter Area | 948 | 953 | 1.005 | 6.72\% | -1358 |
| Year Quarter Area Year*Quarter | 934 | 883 | 0.945 | 6.00\% | -1321 |
| Year Quarter Area Year*Quarter Tqr*Area | 928 | 852 | 0.918 | 2.82\% | -1304 |
| Final model: Year Quarter Area Year*Quarter Tqr*Area Sqr*Area | 926 | 839 | 0.906 | 1.74\% | -1297 |
|  |  |  |  |  |  |
|  |  |  |  |  |  |
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|  |  |  |  |  |  |
|  |  |  |  |  |  |
| Final model factors proportion positives |  |  |  |  |  |
| LR Statistics For Type 3 Analysis |  |  |  |  |  |
| Source | DF | Chi-Square | Pr $>$ ChiSq |  |  |
| quarter | 3 | 204.86 | <. 0001 |  |  |
| area | 1 | 116.41 | <. 0001 |  |  |
| Tqr | 3 | 59.46 | <. 0001 |  |  |
| YEAR | 6 | 46.05 | <. 0001 |  |  |
|  |  |  |  |  |  |
|  |  |  |  |  |  |
| Final model factors positive catch rates |  |  |  |  |  |
| LR Statistics For Type 3 Analysis |  |  |  |  |  |
| Source | DF | Chi-Square | Pr $>$ ChiSq |  |  |
| YEAR | 6 | 15.68 | 0.0156 |  |  |
| quarter | 3 | 67.57 | <. 0001 |  |  |
| area | 1 | 21.99 | <. 0001 |  |  |
| YEAR*quarter | 14 | 91.22 | <. 0001 |  |  |
| Tq**area | 2 | 27.26 | <. 0001 |  |  |
| Sqr*area | 2 | 7.85 | 0.0197 |  |  |
|  |  |  |  |  |  |
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|  |  |  |  |  |  |
| GLM Mixed Model | Neg2LogLike | AIC | BIC |  |  |
|  |  |  |  |  |  |
| Proportion Positives |  |  |  |  |  |
| Quarter Area Tqr Year | 433 | 435 | 438 |  |  |
|  |  |  |  |  |  |
| Positive catch rates |  |  |  |  |  |
| Year Quarter Area Year*Quarter Tqr*Area Sqr*Area | 2549 | 2553 | 2555 |  |  |

Table 2a. Deviance analysis table of explanatory variables in the delta lognormal model for blue shark catch rates (number of sharks per 1000 hooks) from the U.S. pelagic longline fishery observer program (1992-2014). The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.

| Model factors proportion positives | Deviance | Change in deviance | \% of total deviance |
| :---: | :---: | :---: | :---: |
| Null | 4398 |  |  |
| Sqr | 3871 | 527 | 54.33\% |
| Sqr Year | 3712 | 159 | 16.41\% |
| Sqr Year Tqr | 3562 | 149 | 15.40\% |
| Sqr Year Tqr Area | 3490 | 73 | 7.49\% |
| Sqr Year Tqr Area Quarter | 3428 | 62 | 6.37\% |
|  |  |  |  |
|  |  |  |  |
|  |  |  |  |
| Model factors positive catch rates | Deviance | Change in deviance | \% of total deviance |
|  |  |  |  |
| Null | 4558 |  |  |
| Area | 3576 | 983 | 48.12\% |
| Area Quarter | 3184 | 392 | 19.19\% |
| Area Quarter Year | 2998 | 186 | 9.09\% |
| Area Quarter Year Sqr | 2942 | 57 | 2.77\% |
| Area Quarter Year Sqr Tqr | 2893 | 49 | 2.39\% |
| Area Quarter Year Sqr Tqr Year*Quarter | 2688 | 205 | 10.05\% |
| Area Quarter Year Sqr Tqr Year*Quarter Year*Area | 2563 | 125 | 6.11\% |
| Area Quarter Year Sqr Tqr Year*Quarter Year*Area Sqr*Area | 2517 | 46 | 2.27\% |

Table 2b. Deviance analysis table of explanatory variables in the delta lognormal model for blue shark catch rates (number of sharks per 1000 hooks) from the U.S. pelagic longline fishery observer program (2015-2021). The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.

| Model factors proportion positives | Deviance | Change in deviance | \% of total deviance |
| :---: | :---: | :---: | :---: |
| Null | 2201 |  |  |
| Quarter | 2020 | 182 | 44.59\% |
| Quarter Area | 1895 | 125 | 30.81\% |
| Quarter Area Tqr | 1840 | 54 | 13.29\% |
| Quarter Area Tqr Year | 1794 | 46 | 11.30\% |
|  |  |  |  |
|  |  |  |  |
|  |  |  |  |
|  |  |  |  |
| Model factors positive catch rates | Deviance | Change in deviance | \% of total deviance |
|  |  |  |  |
| Null | 1116 |  |  |
| Year | 1060 | 55 | 19.99\% |
| Year Quarter | 1023 | 38 | 13.61\% |
| Year Quarter Area | 953 | 70 | 25.15\% |
| Year Quarter Area Year*Quarter | 883 | 70 | 25.41\% |
| Year Quarter Area Year*Quarter Tqr*Area | 852 | 30 | 10.97\% |
| Year Quarter Area Year*Quarter Tqr*Area Sqr*Area | 839 | 14 | 4.87\% |

Table 3a. Estimates of standardized CPUE (numbers of sharks per 1000 hooks) and coefficients of variation (CV) for blue sharks from the U.S. pelagic longline observer program data (1992-2014). The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.

| Year | Standardized CPUE | CV | Nominal CPUE |
| :---: | :---: | :---: | :---: |
| 1992 | 6.509 | 0.275 | 2.174 |
| 1993 | 10.04 | 0.254 | 2.445 |
| 1994 | 8.375 | 0.254 | 2.222 |
| 1995 | 8.532 | 0.258 | 2.065 |
| 1996 | 6.528 | 0.444 | 1.375 |
| 1997 | 12.53 | 0.289 | 2.483 |
| 1998 | 14.826 | 0.300 | 2.276 |
| 1999 | 6.997 | 0.282 | 2.414 |
| 2000 | 9.037 | 0.273 | 2.464 |
| 2001 | 4.588 | 0.330 | 1.794 |
| 2002 | 5.172 | 0.327 | 1.973 |
| 2003 | 3.619 | 0.302 | 1.176 |
| 2004 | 9.079 | 0.292 | 1.728 |
| 2005 | 3.228 | 0.302 | 1.216 |
| 2006 | 3.651 | 0.300 | 1.472 |
| 2007 | 6.357 | 0.321 | 1.405 |
| 2008 | 6.252 | 0.302 | 1.602 |
| 2009 | 5.961 | 0.301 | 1.738 |
| 2010 | 7.565 | 0.294 | 1.671 |
| 2011 | 13.688 | 0.279 | 2.321 |
| 2012 | 7.229 | 0.287 | 1.645 |
| 2013 | 6.882 | 0.285 | 1.845 |
| 2014 | 6.939 | 0.283 | 1.739 |

Table 3b. Estimates of standardized CPUE (numbers of sharks per 1000 hooks) and coefficients of variation (CV) for blue sharks from the U.S. pelagic longline observer program data (2015-2021). The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.

|  |  |  |  |
| ---: | ---: | ---: | ---: |
| Year | Standardized CPUE | CV | Nominal CPUE |
|  |  |  |  |
| 2015 | 5.196 | 0.286 | 2.095 |
| 2016 | 7.748 | 0.254 | 2.043 |
| 2017 | 6.978 | 0.250 | 1.785 |
| 2018 | 4.581 | 0.299 | 1.653 |
| 2019 | 3.596 | 0.289 | 1.400 |
| 2020 | 3.308 | 0.292 | 1.612 |
| 2021 | 4.081 | 0.308 | 1.857 |

Table 4. Estimates of standardized CPUE (numbers of sharks per 1000 hooks) and coefficients of variation (CV) for blue sharks from the pre-split U.S. pelagic longline observer program data (1992-2021). The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.

| Year | Standardized CPUE | CV | Nominal CPUE |
| ---: | ---: | ---: | ---: |
|  |  |  |  |
| 1992 | 6.109 | 0.270 | 2.174 |
| 1993 | 9.362 | 0.248 | 2.445 |
| 1994 | 8.270 | 0.247 | 2.222 |
| 1995 | 8.215 | 0.252 | 2.065 |
| 1996 | 6.030 | 0.446 | 1.375 |
| 1997 | 12.443 | 0.284 | 2.483 |
| 1998 | 14.726 | 0.293 | 2.276 |
| 1999 | 6.711 | 0.278 | 2.414 |
| 2000 | 9.441 | 0.267 | 2.464 |
| 2001 | 4.877 | 0.324 | 1.794 |
| 2002 | 5.813 | 0.318 | 1.973 |
| 2003 | 3.897 | 0.293 | 1.176 |
| 2004 | 8.941 | 0.285 | 1.728 |
| 2005 | 3.584 | 0.293 | 1.216 |
| 2006 | 3.914 | 0.292 | 1.472 |
| 2007 | 6.665 | 0.312 | 1.405 |
| 2008 | 6.844 | 0.294 | 1.602 |
| 2009 | 6.383 | 0.294 | 1.738 |
| 2010 | 7.451 | 0.286 | 1.671 |
| 2011 | 13.683 | 0.271 | 2.321 |
| 2012 | 7.184 | 0.279 | 1.645 |
| 2013 | 6.864 | 0.278 | 1.845 |
| 2014 | 6.487 | 0.275 | 1.739 |
| 2015 | 6.467 | 0.298 | 2.095 |
| 2016 | 8.442 | 0.274 | 2.043 |
| 2017 | 6.909 | 0.276 | 1.785 |
| 2018 | 4.027 | 0.342 | 1.653 |
| 2019 | 3.664 | 0.306 | 1.400 |
| 2020 | 3.505 | 0.307 | 1.612 |
| 2021 | 3.616 | 0.317 | 1.857 |
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Figure 1. Map of the western North Atlantic Ocean. Areas are as follows: CAR=Caribbean (area 1); GOM=Gulf of Mexico (area 2); FEC=Florida East Coast (area 3); $\mathrm{SAB}=$ South Atlantic Bight (area 4); MAB=Mid-Atlantic Bight (area 5); NEC=Northeast Coastal (area 6); NED=Northeast Distant (area 7); SAR=Sargasso (area 8); NCA=North Central Atlantic (area 9); TUN=Tuna North (area 10); TUS=Tuna South (area 11). Time-area closures (designated by numbers in the map) are as follows: 1- DeSoto Canyon; 2- Florida East Coast; 3- Charleston Bump; 4- Bluefin tuna Northeast Atlantic; 5- Grand Banks; 6- Bluefin tuna spring Gulf of Mexico.

## Blue sharks caught by ICCAT area (observers)



## Blue sharks caught by ICCAT area (observers)



Figure 2. Blue sharks caught by ICCAT area as reported in the pelagic longline observer program for period 1992-2014 (top panel) and period 2015-2021 (bottom panel). Areas are as follows: CAR=Caribbean (area 1); GOM=Gulf of Mexico (area 2); FEC=Florida East Coast (area 3); SAB=South Atlantic Bight (area 4); MAB=MidAtlantic Bight (area 5); NEC=Northeast Coastal (area 6); NED=Northeast Distant (area 7); SAR=Sargasso (area 8); NCA=North Central Atlantic (area 9); TUN=Tuna North (area 10); TUS=Tuna South (area 11).
$\longrightarrow$ Cortes $2016 \longrightarrow$ Standardized P1 $\rightarrow$ Nominal



Figure 3a. Standardized CPUE (sharks/ 1000 hooks) and $95 \%$ confidence intervals for blue shark from the U.S. pelagic longline observer program (P1: 1992-2014) compared to a previous study. All indices are scaled to the mean of the overlapping years (1992-2013) for visualization purposes. The lower panel shows the proportion and number of positive sets by year. The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.


Figure 3b. Standardized CPUE (sharks/ 1000 hooks) and $95 \%$ confidence intervals for blue shark from the U.S. pelagic longline observer program (P2: 2015-2021) compared to standardized CPUE (sharks/ 1000 hooks) from the U.S. pelagic longline observer program (P1: 1992-2014). Even though there is no overlap between these two periods, both indices are scaled to the mean of the years (1992-2013) of index of period 1 for visualization purposes. The lower panel shows the proportion and number of positive sets by year. The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.


Figure 4a. Diagnostic plots of CPUE model from the U.S. pelagic longline observer data (1992-2014) for blue shark. Top: residuals of proportion positive sets; middle: residuals of positive catch; bottom: residual positive catch distribution. The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 20152021 to account for changes in U.S. management regulations.


Figure 4b. Diagnostic plots of CPUE model from the U.S. pelagic longline observer data (2015-2021) for blue shark. Top: residuals of proportion positive sets; middle: residuals of positive catch; bottom: residual positive catch distribution. The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.


Figure 5. Standardized CPUE (sharks/ 1000 hooks) from the U.S. pelagic longline observer program for period 1992-2014 (P1) and period 2015-2021 (P2) compared to the pre-split time series (P1\&P2: 1992-2021). Even though there is no overlap between these two periods, both indices are scaled to the mean of the years (1992-2013) of index of period 1 for visualization purposes. The U.S. pelagic longline observer program data were broken into two periods: 1992-2014 and 2015-2021 to account for changes in U.S. management regulations.


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